TITLE

Meta-Analysis Of Laparoscopic And Open Distal Gastrectomy For Gastric

Carcinoma

AUTHORS

Muhammed Ashraf Memon, FRCS^{1,2,3} Shahjahan Khan. PhD⁴ Rossita Mohamad Yunus, MSc⁴ Richard Barr, MBBS¹ Breda Memon, RGN, LLB⁵

INSTITUTIONS/DEPARTMENTS

¹Department of Surgery, Ipswich Hospital, Queensland, Australia

²Department of Surgery, University of Queensland, Herston, Queensland, Australia

³Faculty of Health Sciences and Medicine, Bond University, Gold Coast, Queensland,

Australia

⁴Department of Mathematics and Computing, Australian Centre for Sustainable

Catchments, University of Southern Queensland, Toowoomba, Queensland, Australia

⁵Health & Social Care Department, Higher Education Building, Blackburn College,

Blackburn, Lancashire, United Kingdom

REPRINTS/CORRESPONDENCE

M. A. Memon, FRCS, Ipswich Hospital, Chelmsford Avenue, Ipswich, Queensland,

Australia

Tel: +61 448614170 Fax: +61 7 38101592

Email: mmemon@yahoo.com

SECTION OF THE JOURNAL

Meta-analysis

KEY WORDS

Gastrectomy; Laparoscopic method, Comparative studies; Meta-analysis;

Randomised controlled trials; Patient's outcome; Intraoperative complications;

Postoperative complications; Hospitalization; Human

ABSTRACT:

Objectives: The aim was to conduct a meta-analysis of the randomized evidence to determine the relative merits of laparoscopic assisted (LADG) and open (ODG) distal gastrectomy for proven gastric cancer.

Data sources and review methods: A search of the Medline, Embase, Science Citation Index, Current Contents and PubMed databases identified all randomized clinical trials (RCTs) that compared LADG and OGD and were published in the English language between January 1990 and the end of June 2007. The meta-analysis was prepared in accordance with the Quality of Reporting of Meta-analyses (QUOROM) statement. The eight outcome variables analysed were operating time, blood loss, retrieval of lymph nodes, oral intake, hospital stay, postoperative complications, tumour recurrence and mortality. Random effects meta-analyses were performed using odds ratios (OR) and weighted mean differences (WMD).

Results: Four trials were considered suitable for meta-analysis. A total of 81 patients underwent LADG and 80 had ODG. For only one of the eight outcomes, the summary point estimates favoured LADG over ODG; there was a significant reduction of 104.26 mls in the intra-operative blood loss for LADG (WMD, -104.26, 95% confidence interval (c.i.) -189.01 to -19.51; P = 0.0159). There was however a 83.08 minutes longer duration of operating time for the LADG group compared with the ODG group (WMD 83.08, 95% c.i. 40.53 – 125.64; P = 0.0001) and significant reduction in lymph nodes harvesting of 4.34 lymph nodes in the LADG group (WMD -4.3, 95% c.i. -6.66 to -2.02; P = 0.0002). Other outcome variables such as time to commencement of oral intake (WMD -0.97, 95% c.i. -2.47 to 0.54 ; P = 0.2068);

duration of hospital stay (WMD -3.32, 95% c.i. -7.69 to 1.05; P = 0.1365); rate of complications (OR 0.66, 95% c.i. 0.27 to 1.60; P = 0.3530); mortality rates (OR 0.94, 95% c.i. 0.21 – 4.19; P = 0.9363) and tumour recurrence (OR 1.08 (95% c.i. 0.42 – 2.79; P = 0.8806) were not found to be statistically significant for either group. However, for commencement of oral intake, duration of hospital stay and complication rate, the trend was in favour of LADG.

Conclusion: LADG was associated with a significantly reduced rate of intraoperative blood loss, at the expense of significantly longer operating time and significantly reduced lymph node retrieval compared to its open counterpart. Mortality and tumour recurrence rates were similar between the two groups. Furthermore, time to commencement of oral intake, post procedural discharge from hospital and the perioperative complication rate although not significantly different between the two groups did suggest a positive trend toward LADG. Based on this meta-analysis, the authors cannot recommend the routine use of LADG over ODG for the treatment of distal gastric cancer. However, significant limitations exist in the interpretation of this data due to the limited number of published randomised control trials, the small sample sizes to date, and the limited duration of follow up. Further large multicentre randomised controlled trials are required to delineate significantly quantifiable differences between the two groups.

INTRODUCTION:

Minimal access gastrointestinal surgery for gastric cancer in the form of laparoscopic distal gastrectomy was introduced 13 years ago by the Japanese surgeons¹. However, its wider acceptance as an alternative to an open approach remains a contentious issue. There could be a number of reasons for such a slow acceptance which include the complexity of the procedure especially the subsequent reconstruction of the alimentary tract and oncological adequacy and safety in terms of R0 resection, lymph nodes harvesting and tumour recurrence at the trocar sites². The other factor which may have slowed its progress is the extent of the associated lymph node dissection, an issue which is viewed differently by the Japanese and Western Surgeons³⁻⁸. However, recent years have also seen a tremendous amount of advancement in the development of laparoscopic instruments which, coupled with increasing experience in the performance of complex laparoscopic gastrointestinal procedures, have led to the expansion of minimal access surgery for both benign and malignant gastric procedures.

Our aim therefore was to conduct an appraisal, based on a meta-analysis of pooled data from 4 available randomized clinical trials⁹⁻¹², to compare the laparoscopic and open methods of distal gastrectomy for gastric cancer with a view to providing some clarity on a number of issues considered to be controversial. This meta-analysis was prepared in accordance with the Quality of Reporting of Meta-analyses (QUOROM) statement¹³.

MATERIAL AND METHODS:

Randomized clinical trials of any size that compared LADG with ODG for gastric cancer, and were published in full in peer-reviewed journals in the English language between January 1990 and the end of June 2000, were included. Unpublished studies and abstracts presented at national and international meetings were excluded. Trials were identified by conducting a comprehensive search of Medline, Embase, Science Citation Index, Current Contents and PubMed databases, using medical subject headings (MESH) 'gastrectomy', 'comparative study', 'prospective studies', 'randomized controlled trials', 'random allocation' and 'clinical trial'. A manual search of the bibliographies of relevant papers was also carried out to identify trials for possible inclusion. Data extraction and critical appraisal were carried out by three authors independently (BM, RB, MAM). Eight outcome variables were considered most suitable for analysis: operating time, blood loss, retrieval of lymph nodes, oral intake, hospital stay, postoperative complications, tumour recurrence and mortality rate. The quality of the randomized clinical trials was assessed using Jadad's scoring system¹⁴ by the two authors (BM, RB) (Table 1).

STATISTICAL ANALYSIS:

Meta-analyses were performed using odds ratios (ORs) for binary and weighted mean differences (WMDs) for continuous outcome measures. The slightly amended estimator¹⁵ of OR is used to avoid the computation of reciprocal of zeros among observed values in the calculation of the original OR. Random effects models by using the weighted method approach were used to combine the data and statistical heterogeneity was assessed using the χ^2 test¹⁶. A sensitivity analysis was carried out to assess the impact of study quality on the results, by excluding poor-quality studies

(Jadad score 1). Funnel plots (Fig 9) were constructed to detect publication bias in meta-analysis by plotting both size and precision (1/standard error) against the treatment effect (MDs/LORs) for each outcome variables¹⁶⁻¹⁸. Sixteen funnel plots are plotted for 8 outcome variables. Eight of them are size based funnel plots while another 8 are standard error based funnel plots. For this meta-analysis, the number of points in the funnel plots (the number of studies) are few i.e. 4 in each plot, therefore the detection of bias is limited^{17,19}. The funnel plots were produced to show (a) that the conclusion about the shape or visual interpretation may alter by plotting treatment effects against precision error instead of the sample size^{16,18} and (b) the limitation in the use of funnel plots to detect publication bias when the number of studies are small. All estimates were obtained using a computer program written by R, an open source software (copyright © 1998–2007 by Kurt Hornik)²⁰. All plots were obtained using a computer program written by R, an open source software (sufficient to be the sum of software (sufficient to be the sum of sufficient to be the sum of sufficient to be the sum of sufficient to be sufficient to b

RESULTS:

A total of 4 randomized prospective clinical trials that included 161 distal gastrectomies (LADG 81, ODG 80) were considered suitable for meta-analysis. In general, the quality of the studies was poor on critical appraisal (mean quality score 2.7 of 5) (Table 1). This was because the method of randomization was not defined in every study, it was not possible to blind study participants and investigators for these procedures, and a description of withdrawals and drop-outs was not always provided. This is not an uncommon feature amongst the surgical RCTs, and has been observed in many reviews and meta-analyses of surgical trials²¹⁻²³. The pooled data (OR and WMD) for the eight outcomes are summarized in Table 2 and Figs 1–8. As

statistically significant heterogeneity was evident for the majority of outcome variables, random effects models were used to combine the data²⁴. A total of 81 patients underwent LADG and 80 had ODG. For only one of the eight outcome variables, the summary point estimates favoured LADG over ODG; there was a significant reduction of 104.26 mls in the intra-operative blood loss for LADG (WMD -104.26, 95% confidence interval (c.i.) -189.01 to -19.51; P = 0.0159). There was however a 83.08 minutes longer duration of operating time for the LADG group compared with the ODG group (WMD 83.08, 95% c.i. 40.53 - 125.64; P = 0.0001) and significant reduction in lymph nodes harvesting of 4.34 lymph nodes in the LADG group (WMD -4.3, 95% c.i. -6.66 to -2.02; P = 0.0002). Other outcome variables such as time to commencement of oral intake (WMD -0.97, 95% c.i. -2.47 to 0.54; P = 2068); duration of hospital stay for LADG (WMD -3.32, 95% c.i. -7.69 to 1.05; P = 0.1365); rate of complications (OR 0.66, 95% c.i. 0.27 to 1.60; P = 0.3530); mortality rates (OR 0.94, 95% c.i. 0.21 - 4.19; P = 0.9363) and tumour recurrence (OR 1.08, 95% c.i. 0.42 - 2.79; P = 0.8806) were not found to be statistically significant for either group. However, for commencement of oral intake, duration of hospital stay and complication rate, the trend was in favour of LADG.

DISCUSSIONS:

The proponents of LADG argue that the procedure is superior to ODG because it is associated with less postoperative pain, reduced peri-operative blood loss, quicker return to gastrointestinal function, faster hospital discharge, an earlier return to work and unrestricted physical activity, and a better cosmetic result². The opponents, however, argue that there is a higher incidence of major intraoperative and postoperative complications because of the complexity of the procedure and absence

of tactile sensation, substantially greater costs, much longer anaesthetic and operating time, decreased numbers of lymph nodes harvesting essential for oncological adequacy, insufficient surgical resection margins and potential for cancer reimplantation at trocar sites². Furthermore, long-term consequences are unknown. This controversy has encouraged a number of investigators⁹⁻¹² to initiate randomized clinical trials in an attempt to address some of these issues. However all these comparative trials have recruited a limited number of patients to date and the long term follow-up is not available. To clarify some of these issues, we undertook the present meta-analysis, concentrating on eight treatment variables that could be analysed objectively. To our knowledge no meta-analyses or a systematic review on this subject has been undertaken or published.

All the trials reported the duration of operation. The meta-analysis revealed statistically significant longer operating time for LADG than for ODG (Fig. 1). This has important implications for both patients and the healthcare providers. Longer operations expose patients to a protracted anaesthesia, which may increase the morbidity and even mortality rates especially in older patients with co-morbidities. The vast majority of patients in these four RCTs were in their late 50s or early 60s with concomitant co-morbidies including cardio-pulmonary issues, diabetes etc. Longer operating and anaesthesia times also increase the direct cost of the procedure. Even with experience, the operating time for LADG has remained substantially longer to date. Without exception, all the RCTs (Fig 1) have clearly shown longer operating time for LADG by the authors who are considered experienced upper GI and laparoscopic surgeons. The longer operating time for LADG may in part reflect an early learning curve, as this is a relatively new procedure. Furthermore, the operating

time for LADG also includes the time for setting up laparoscopic equipments. Other reasons include lack of tactile sensation, the complexity of procedure, the postresectional reconstruction of gastrointestinal tract and the extent of lymphadenectomy performed at least by the Japenese surgeons.

Regarding the intra-operative blood loss, three out of four RCTs^{9,10,12} (Fig 2) have shown decreased blood loss for LADG compared to ODG. This translates into decreased transfusion requirement in the peri-operative period with its inherent risk of acute or late adverse effects such as acute lung injury, volume overload, hypothermia, graft versus host disease and immunomodulatory effects to name but a few. The last side effect is especially important in cancer patients as a number of studies and a meta-analysis has suggested a significant deleterious transfusion effect in all cancer sites, except for breast^{25,26}. Furthermore, the blood transfusion economics has not been addressed in any of these studies. A multicentre study on blood transfusion cost performed in 1991²⁷ revealed the average hospital cost per unit transfused was \$155 which would be far higher now. It is therefore evident from this meta-analysis that LADG have biological, immunological and economical benefits for the patient and the health care system by reducing peri-operative transfusion needs.

Concerning oncological adequacy for lymph node harvesting, the number of lymph nodes retrieved laparoscopically in all these RCTs were sufficient (Fig 3) where the global standard for adequate staging is concerned, emphasizing the oncological capability of laparoscopic gastric procedures². In fact, in none of the RCTs the lymph nodes retrieval for the two procedures has shown any significant statistical difference. However, when the results were pooled together (Fig 3), there was a statistically

significant reduction in lymph node harvesting for LADG compared to ODG which may translate into an overall survival disadvantage for patients having LADG. As the long term results for the majority of these trials have not been published, this assumption is difficult to corroborate.

The debate concerning the merits and risks of extended lymph node clearance during gastrectomy for cancer remains a contentious issue. A number of authors still feel that clinical benefit from extended lymphadenectomy for gastric cancer has no proven benefit and may even be counterproductive. A large retrospective study from Finland²⁸ analysing 223 patients (D1=114, D2=109) undergoing curative gastrectomy found the surgical complication to be statistically higher for D2 cohort (33% vs 16.8%) although hospital mortality was similar between the two groups. Furthermore D2 lymphadenectomy was associated with a longer postoperative hospital stay, operating time, blood loss and increased need for blood transfusions. A large Japanese multicentre RCT²⁹ consisting of 523 patients once again showed higher morbidity for the extended (D2) surgery compared to the D1 group (28.1% vs 20.9%) for curative gastric cancer. This difference however did not reach statistical significance. Nonetheless, the authors felt that extended (D2) surgery can be added without increasing major surgical complications in this setting. Yet another RCT from The Netherlands³⁰ analysing 711 patients (D1=380, D2=331) has shown a significantly higher morbidity and mortality for the D2 group (25% vs 43% and 4% vs 10% respectively) without affecting the five year survival rate. The Cochrane Review³¹ has shown no survival benefit for extended lymph node dissection but showed increased postoperative mortality and morbidity. Miura et al³² performed a critical reappraisal from the viewpoint of lymph node retrieval and found that laparoscopic D2 resection

harvested sufficient numbers of nodes for adequate TNM classification in 86% of cases. However, a significantly greater number of lymph nodes were harvested by the open method. They concluded that the extent of lymphadenectomy achieved by current laparoscopic procedures approaches the global standard for accurate staging, although performing extended resection laparoscopically as recommended in Japan remains a challenge and is a time consuming process. The authors therefore suggested that laparoscopic gastrectomy is only justified for more advanced disease under the setting of clinical trials.

Except for Kitano et al⁹, all the other RCTs¹⁰⁻¹² showed early resumption of oral intake by the patients undergoing LADG (Fig 4). The pooled data showed a positive trend for LADG, however this did not reach statistical significance. Three out of the four trials^{9,11,12} which reported on postoperative recovery of gastrointestinal function (passage of first flatus) showed that patients in all these trials have a quicker return of their gastrointestinal function and in two of these trials, this difference was significant^{9,12}. Quicker return of gastrointestinal function has a direct impact on early resumption of diet which is shown in these trials and this allows early discharge with economical benefits.

Three out of four RCTs¹⁰⁻¹² showed a trend towards earlier discharge from hospital after LADG (Fig 5). Pooling the data from these trials failed to show any difference in the discharge data for these two procedures although the trend favoured LADG (Table 2). Early-discharge is associated with lower medical direct, non-medical direct, and indirect costs than conventional inpatient care. Cost savings per patient therefore can be significant. Furthermore, early discharge also has a positive effect on pressure on

hospital beds which in certain countries have decreased due to restructuring of health services and have a direct impact on elective admissions. It is entirely possible that larger RCTs may show that LADG indeed have a clear cut advantage over ODG in terms of hospital discharge. We eagerly await any such data.

None of the present trials have provided comparative data on patients returned to normal activity following LADG and ODG. However, all the RCTs have shown that the frequency, dose and duration of analgesia requirement for LADG have been lower in the peri-operative period. This is most likely due to the absence of a large abdominal incision in LADG. Two of the four RCTs^{9,10} have shown significantly early ambulation in their patients undergoing LADG. All these findings translate into a quicker return of biological functions, early hospital discharge and quicker return to normal activities. Obviously an objective assessment would be ideal using one of the health profile questionnaires which measures physical, mental, or emotional problems or limitations in patients' daily life in the immediate and late peri-operative periods. This may have a major repercussion both for the employers and society in general.

As far as the complication rate of these two procedures is concerned, the present analysis showed a higher incidence of peri-operative complications after ODG, (Fig. 6). However this did not reach statistical significance when compared to LADG. Because laparoscopic surgery avoids a large abdominal incision, this decreases the incidence of postoperative pain which in turn decreases the incidence of atelectasis, hypoventilation, pneumonia and coronary ischaemia. A number of RCTs and observational studies have shown that laparoscopic procedures are associated with less suppression of FVC and FEV₁ compared to their open counterpart³³⁻³⁶. In all the

RCTs of LADG vs ODG, the authors have observed more cardiorespiratory complications following ODG compared to LADG. Also the incidence of wound infection is higher for the open cohort because of the larger incision size. It is well known that reducing the number of complications should produce significant savings with an equal or better health outcome. The laparoscopic gastric procedure in this meta-analysis has shown 34% reduction in the relative odds of complications which although not statistically significant certainly translate into better outcome for the patient and the health care system.

Lastly, there was no significant difference in mortality rate (Fig 7) and tumour recurrence (Fig 8) between the two procedures. There could be a number of explanations for such parity. First of all the number of patients in all these RCTs are relatively small which may have masked the true difference in mortality rate. Secondly, the follow-up data is short and therefore the real difference in the tumour recurrence may not be apparent presently. Once the data from these trials is matured, one will be able to get a clearer picture on these outcomes variables.

CONCLUSIONS:

The present meta-analysis included a total of 161 distal gastrectomies for cancer, the largest body of information so far available for the comparison of LADG and ODG in the English language literature. Laparoscopic gastrectomy was associated with significantly decreased blood loss and positive trends towards fewer postoperative complications, quicker commencement of oral intake, earlier hospital discharge and early mobilization with decreased requirement for analgesia, but at the expense of a significantly longer operating time and fewer lymph nodes retrieval. Based on these data the authors feel the clear cut benefits of LADG over ODG are rather limited and its widespread adaptation cannot be recommended. However, significant limitations exist in the interpretation of this data due to the limited number of published randomised control trials, the small sample sizes to date, and the limited duration of follow up. Further large multicentre randomised controlled trials are required to delineate significantly quantifiable differences between the two groups. Nonetheless, it may be concluded that LADG is a safe and effective alternative to ODG and is justifiable under the setting of clinical trials.

ACKNOWLEDGMENT:

Authors' contributions:

MAM was responsible for the concept and design of this meta-analysis. Furthermore he takes responsibility for the integrity of the work as a whole, from inception to published article.

MAM, RB and BM were responsible for acquisition and interpretation of the data. SK and RMY were involved in analysing and interpretation of the data in depth from the statistical point of view.

All authors were involved in drafting the manuscript and revising it critically for important intellectual content and have given final approval of the version to be published. Furthermore all authors have participated sufficiently in the work to take public responsibility for its content

Conflicts of interests:

None

Sources of funding and support:

None

Sponsor(s):

None

Table 1: Jadad's score

Authors	Year	Country	Randomisation	Blinding	Withdrawal	Total Jadad's Score
Kitano et al ⁹	2002	Japan	2	0	1	3
Lee et al ¹¹	2005	Korea	2	0	1	3
Hayashi et al ¹²	2005	Japan	2	0	1	3
Huscher et al ¹⁰	2005	Italy	1	0	1	2

Table 2 Summary of pooled data comparing LADG and ODG

Variables	Pooled OR or WMD	Test for o	verall effect	Test for heterogeneity	
		Ζ	р	χ^{2}	р
Duration of operating time (min)	83.08 (40.53, 125.64) ‡	3.83	0.0001	91.93	< 0.0001
Intraoperative blood loss (mls)	-104.26 (-189.01, -19.51) ‡	-2.41	0.0159	13.56	0.0037
Lymph nodes harvesting	-4.34 (-6.66, -2.02) ‡	-3.66	0.0002	1.68	0.6421
Time to commencement of oral intake (day)	-0.97 (-2.47, 0.54) ‡	-1.26	0.2068	36.90	< 0.0001
Duration of hospital stay (day)	-3.32 (-7.69, 1.05) ‡	-1.49	0.1365	33.54	< 0.0001
Complication rate	0.66 (0.27, 1.60) *	-0.93	0.3530	4.87	0.1819
Mortality rate	0.94 (0.21, 4.19) *	-0.08	0.9363	0.78	0.8547
Tumour recurrence	1.08 (0.42, 2.79) *	0.15	0.8806	0.01	0.9998

Values in parentheses are 95 percent confidence intervals. *OR odds ratio; ‡ WMD weighted mean difference.

From the test of heterogeneity, 4 variables (duration of operating time (min), intraoperative blood loss (mls), time to commencement of oral intake (day) and duration of hospital stay (day) are found rejecting the null hypothesis that is all true treatment effects for 4 studies (references) are the same (statistically significant with p-value <0.004.

The random effect method is applied to all variables to find the pooled point estimate and confidence interval of the mean effect.

Later the overall effect test is used to test whether the pooled estimate is different from zero (i.e. whether the difference due to LADG and ODG is statistically significant). The test of overall effect shows statistically significant longer operating time for LADG than ODG (p-value =0.0001) but statistically significant more blood loss and more lymph nodes for ODG than LADG (p-value <0.05). The data fails to reject the null hypothesis that LADG and ODG are not different for time to commencement of oral intake, duration of hospital stay, complication rate, mortality rate and tumour recurrence.

Mean of duration of operation in minutes (standard deviation)

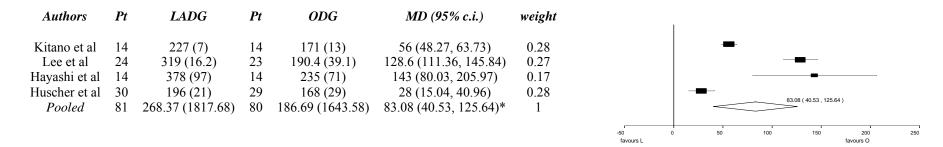


Fig. 1 Values in left panel are mean (standard deviation), mean difference (95% c.i.), weighted mean difference (95% c.i.) and weight. In the graph, squares indicate point estimates of treatment effect (mean difference, i.e. mean for LADG group of patients – mean for ODG group of patients) with the size of the squares representing the weight attribute to each study. The horizontal lines represent 95% confidence interval for means differences. The pooled estimate of operating time (minutes) is the weighted mean difference, obtained by combining all means differences using the inverse weighted method and is represented by the diamond and the size of the diamond depicts the ninety-five percent confidence interval. Values to the left of the vertical line at zero favour LADG.

Mean of blood loss in mls (standard deviation)

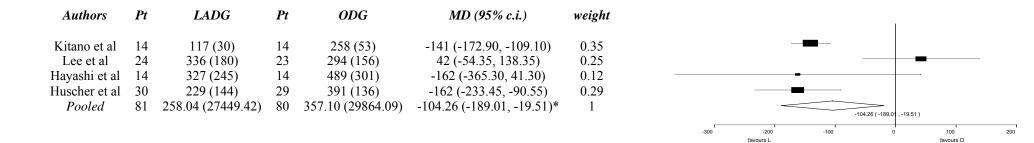


Fig. 2 Values in left panel are mean (standard deviation), mean difference (95% c.i.), weighted mean difference (95% c.i.) and weight. In the graph, squares indicate point estimates of treatment effect (mean difference, i.e. mean for LADG group of patients – mean for ODG group of patients) with the size of the squares representing the weight attribute to each study. The horizontal lines represent 95% confidence interval for means differences. The pooled estimate of blood loss (mls) is the weighted mean difference, obtained by combining all means differences using the inverse weighted method and is represented by the diamond and the size of the diamond depicts the ninety-five percent confidence interval. Values to the left of the vertical line at zero favour LADG.

Mean of lymph nodes harvested (standard deviation)

Authors	Pt	LADG	P t	ODG	MD (95% c.i.)	weight
Kitano et al	14	20.2(3.6)	14	24.9 (3.5)	-4.7(-7.33, -2.07)	0.78
Lee et al	24	31.8(13.5)	23	38.1(15.9)	-6.3(-14.75,2.15)	0.08
Hayashi et al	14	28(14)	14	27(10)	1(-8.01,10.01)	0.07
Huscher et al	30	30(14.9)	29	33.4(17.4)	-3.4(-11.68,4.88)	0.08
Pooled	81	28.51(171.11)	80	32.14 (203.93)	-4.3(-6.66, -2.02)*	1

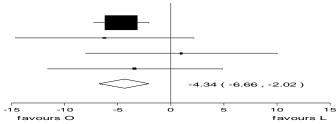


Fig 3 Values in left panel are mean (standard deviation), mean difference (95% c.i.), weighted mean difference (95% c.i.) and weight. In the graph, squares indicate point estimates of treatment effect (mean difference, i.e. mean for LADG group of patients – mean for ODG group of patients) with the size of the squares representing the weight attribute to each study. The horizontal lines represent 95% confidence interval for means differences. The pooled estimate of lymph nodes harvested (units) is the weighted mean difference, obtained by combining all means differences using the inverse weighted method and is represented by the diamond and the size of the diamond depicts the ninety-five percent confidence interval. Values to the left of the vertical line at zero favour ODG.

Mean of oral intake in days (standard deviation)

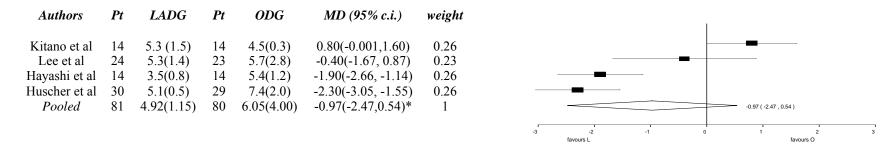


Fig. 4 Values in left panel are mean (standard deviation), mean difference (95% c.i.), weighted mean difference (95% c.i.) and weight. In the graph, squares indicate point estimates of treatment effect (mean difference, i.e. mean for LADG group of patients – mean for ODG group of patients) with the size of the squares representing the weight attribute to each study. The horizontal lines represent 95% confidence interval for means differences. The pooled estimate of time to commencement of oral intake (days) is the weighted mean difference, obtained by combining all means differences using the inverse weighted method and is represented by the diamond and the size of the diamond depicts the ninety-five percent confidence interval. Values to the left of the vertical line at zero favour LADG.

Mean of duration of hospital stay in days (standard deviation)

.

Authors	Pt	LADG	Pt	ODG	MD (95% c.i.)	weight
Kitano et al	14 24	17.6 (2.6)	14 23	16(0.4) 17.3(15.5)	1.6(0.22, 2.98) -6.1(-12.65,0.45)	0.29 0.18
Lee et al Hayashi et al	24 14	11.2(4.2) 12(2)	23 14	18(6)	-6(-9.31,-2.69)	0.18
Huscher et al <i>Pooled</i>	30 81	10.3(3.6) 12.1(11.81)	29 80	14.5(4.6) 16.18(83.53)	-4.2(-6.32, -2.09) -3.32(-7.69, 1.05)*	0.28 1

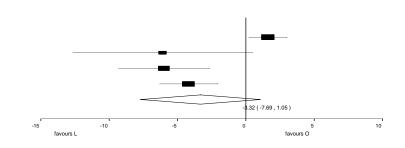
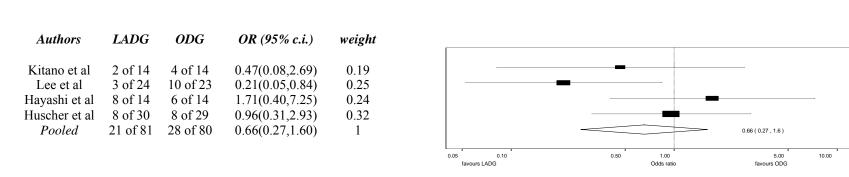


Fig. 5 Values in left panel are mean (standard deviation), mean difference (95% c.i.), weighted mean difference (95% c.i.) and weight. In the graph, squares indicate point estimates of treatment effect (mean difference, i.e. mean for LADG group of patients – mean for ODG group of patients) with the size of the squares representing the weight attribute to each study. The horizontal lines represent 95% confidence interval for means differences. The pooled estimate of duration of hospital stay (days) is the weighted mean difference, obtained by combining all means differences using the inverse weighted method and is represented by the diamond and the size of the diamond depicts the ninety-five percent confidence interval. Values to the left of the vertical line at zero favour LADG.



Odds ratio for complications

Fig. 6 In the graph, squares indicate point estimates of treatment effect (odds ratio for LADG over ODG groups) with the size of the squares representing the weight attribute to each study. The horizontal lines represent 95% confidence interval for odds ratio. The pooled estimate for complication rate (is the pooled odds ratio obtained by combining all odds ratio of the four studies using the inverse weighted method) and is represented by the diamond and the size of the diamond depicts the ninety-five percent confidence interval. Values to the left of the vertical line at one favour LADG.

Odds ratio for mortality

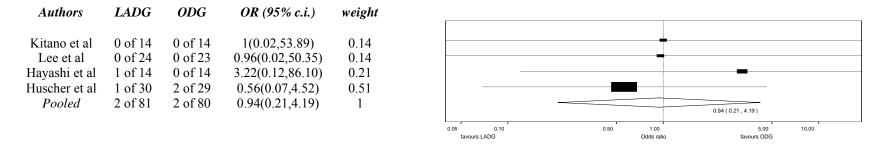


Fig. 7 In the graph, squares indicate point estimates of treatment effect (odds ratio for LADG over ODG groups) with the size of the squares representing the weight attribute to each study. The horizontal lines represent 95% confidence interval for odds ratio. The pooled estimate for mortality rate (is the pooled odds ratio obtained by combining all odds ratio of the four studies using the inverse weighted method) and is represented by the diamond and the size of the diamond depicts the ninety-five percent confidence interval. Values to the left of the vertical line at one favour LADG.

Odds ratio for tumour recurrence

Authors	LADG	ODG	OR (95% c.i.)	weight
Kitano et al	0 of 14	0 of 14	1.00(0.02,53.89)	0.06
Lee et al	0 of 24	0 of 23	0.96(0.02,50.35)	0.06
Hayashi et al	0 of 14	0 of 14	1.00(0.02,53.90)	0.06
Huscher et al	11 of 30	10 of 29	1.10(0.38,3.12)	0.83
Pooled	11 of 81	10 of 80	1.08(0.42,2.79)	1

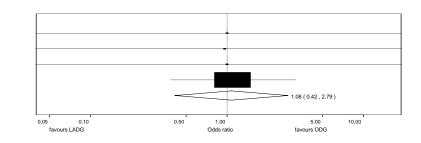


Fig. 8 In the graph, squares indicate point estimates of treatment effect (odds ratio for LADG over ODG groups) with the size of the squares representing the weight attribute to each study. The horizontal lines represent 95% confidence interval for odds ratio. The pooled estimate for tumour recurrence (is the pooled odds ratio obtained by combining all odds ratio of the four studies using the inverse weighted method) and is represented by the diamond and the size of the diamond depicts the ninety-five percent confidence interval. Values to the left of the vertical line at one favour LADG.

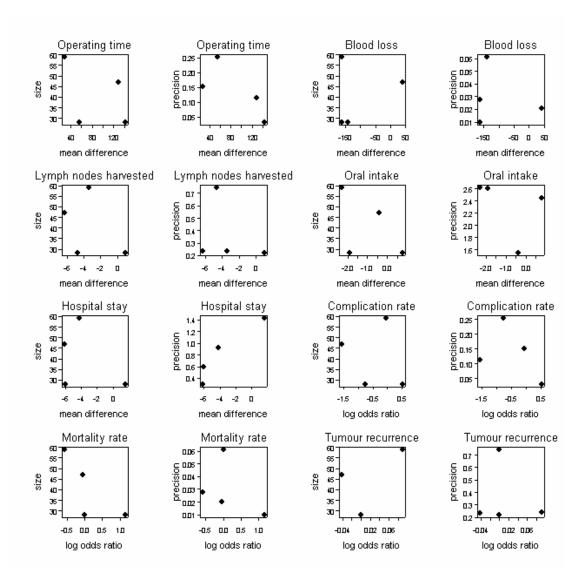


Fig 9: Funnel plots for every treatment from the studies included in meta-analysis. Precision = 1/standard error.

REFERENCES:

1. Kitano S, Iso Y, Moriyama M. Laparoscopy- assisted Billroth I gastrectomy. Surg Laparosc Endosc 1994; 4:146-8.

 Shehzad K, Mohiuddin K, Nizami S, Sharma H, Khan IM, Memon B, Memon MA.
 Current status of minimal access surgery for gastric cancer. Surg Oncol 2007; 16: 85-98.

3. Kajitani T. The general rules for the gastric cancer study in surgery and pathology. Japanese research for gastric cancer. Jap J Surg 1981; 11:127-139.

4. Japanese Gastric Cancer Association. The guidelines for the treatment of gastric cancer. Tokyo: Kachara Co. 2001.

5. Maruyama K, Gunven P, Okabayashi K. Lymph node metastases of gastric cancer: general pattern in 1931 patients. Ann Surg 1987; 210: 596-602.

6. Maruyama K, Okabayashi K, Kinoshita T. Progress in gastric cancer surgery in Japan in Japan and its limits of radicality. Word J Surg 1987; 11: 418-425.

7. Noguchi Y, Imada T, Matsumoto A. Radical surgery for gastric cancer: a review of the Japanese experience. Cancer 1986; 64: 2053-2062.

8. Cuschieri A., Weeden S, Fieldig J. Patient survival after D1 and D2 resections for gastric cancer: long term results of MRC randomized surgical trials. Br J Cancer 1999; 79: 1522-1530.

9. Kitano S, Shiraishi N, Fujii K, Yasuda K, Inomata M, Adachi Y. A randomized controlled trial comparing open vs laparoscopy-assisted distal gastrectomy for the treatment of early gastric cancer: an interim report. Surgery 2002; 131:S306-11.

10. Huscher CGS, Mingoli A, Sgarzini G, Sansonetti A, Di Paola M, Rcher A,Ponzano C. Laparoscopic versus open subtotal gastrectomy for distal gastric cancer:five year results of a randomized prospective trial. Ann Surg 2005; 241:232-7.

11. Lee JH, Han HS, Lee JH. A prospective randomized study comparing open vs laparoscopy-assisted distal gastrectomy in early gastric cancer: early results. Surg Endosc 2005; 19:168-73.

12. <u>Hayashi H, Ochiai T, Shimada H, Gunji Y.</u> Prospective randomized study of open versus laparoscopy-assisted distal gastrectomy with extraperigastric lymph node dissection for early gastric cancer. Surg Endosc 2005; 19:1172-6.

13. Moher D, Cook DJ, Eastwood S, Olkin I, Rennie D, Stroup DF. Improving the quality of reports of meta-analyses of randomised controlled trials: the QUOROM statement. Quality of Reporting of Meta-analyses. Lancet 1999; 354: 1896–1900.

14. Jadad AR, Moore RA, Carroll D, Jenkinson C, Reynolds DJ, Gavaghan DJ.Assessing the quality of reports of randomized clinical trials: is blinding necessary?Control Clin Trials 1996; 17: 1–12.

Agresti A. Categorical Data Analysis (2nd ed). Wiley & Sons Inc: Canada 2002 pp
 1-734.

16. Sutton AJ, Abrams KR, Jones DR, Sheldon TA, Song F. Methods for Meta-Analysis in Medical Research. John Wiley & Sons Inc: England 2000 pp 1-346.

17. Egger M, Smith GD, Schneider M, Minder C. Bias in meta-analysis detected by a simple, graphical test. Br Med J 1997; 315: 629-634.

 Tang JL, Liu JLY. Misleading funnel plot for detection of bias in meta-analysis. J Clin Epidemiol 2000; 53: 477-484.

 Span J, Carière E, Croockewitt S, Smits P. Publication bias, effects on the assessment of Rosiglitasone. Proc Dutch Soc Clin Pharmacol Biopharm Meet, April 2006. Br J Clin Pharmacol 2006; 62: 732.

20. Hornik K. The R **FAQ.** Version 2.6.2007-10-22. ISBN 3-900051-08-9. http://cran.r-project.org/doc/FAQ/R-FAQ.html

21. Solomon MJ, Laxamana A, Devore L, McLeod RS. Randomized controlled trials in surgery. Surgery 1994; 115: 707–712.

22. McLeod RS, Wright JG, Solomon MJ, Hu X, Walters BC. Lossing A.Randomized controlled trials in surgery: issues and problems. Surgery 1996; 119: 483–486.

23. Horton R. Surgical research or comic opera: questions, but few answers. Lancet 1996; 347: 984–985.

24. Sutton AJ, Abrams KR, Jones DR, Sheldon TA, Song F. Methods for Metaanalysis in Medical Research. JohnWiley: Chichester, 2000.

25. Dellinger EP, Anaya DA. Infectious and immunologic consequences of blood transfusion. Crit Care 2004; 8 Suppl 2: S18-23.

26. <u>Vamvakas EC</u>. Perioperative blood transfusion and cancer recurrence: metaanalysis for explanation. Transfusion 1995; 35: 760-8.

27. Forbes JM, Anderson MD, Anderson GF, Bleecker GC, Rossi EC, Moss GS.Blood transfusion costs: a multicenter study. Transfusion 1991; 31: 318-323.

Danielson H, Kokkola A, Kiviluoto T, Siren J, Louhimo J, Kivilaakso E,
 Puolakkainen P. Clinical outcome afer D1 vs D2-3 gastrectomy for treatment of
 gastric cancer. Scand J Surg 2007; 96: 35-40.

29. Sano T, Sasako M, Yamamoto S, Nashimoto A, Kurita A, Hiratsuka M, Tsujinaka T, Kinoshita T, Arai K, Yamamura Y, Okajima K. Gastric cancer surgery: morbidity and mortality resuls from a prospective randomized controlled trial comparing D2 and extended paraaortic lymphadenectomy – Japan Clinical Oncology Group study 9501.

30. Bonenkamp JJ, Hermans J, Sasako M, van de Velde CJ, Welvaart K, Songun I, Meyer S, Plukker JT, Van Elk P, Obertp H, Gouma DJ, van Lanschot JJ, Taat CW, de Graaf PW, von Meyenfeldt MF, Tilanus H; Dutch Gastric Cancer Group. N Engl J Med 1999; 340: 908-914.

31. McCulloch P, Nita ME, Kazi H, Gama-Rodrigues J. Extended versus limited lymph nodes dissection technique for adenocarcinoma of the stomach. Cochrane Database Syst Rev 2004; 18: CD001964.

32. <u>Miura S, Kodera Y, Fujiwara M, Ito S, Mochizuki Y, Yamamura Y, Hibi K, Ito K,</u> <u>Akiyama S, Nakao A.</u> Laparoscopy-assisted distal gastrectomy with systemic lymph node dissection: a critical reappraisal from the viewpoint of lymph node retrieval. *J* Am Coll Surg 2004; 198:933-8.

33. Schauer PR, Luna J, Ghiatas AA, Glen ME, Warren JM, Sirinek KR. Pulmonary function after laparoscopic cholecystectomy. Surgery 1993; 114:389-99.

34. McMahon AJ, Baxter JN, Kenney G, O'Dwyer PJ. Ventilatory and blood gas change during laparoscopic and open cholecystectomy. Br J Surg 1993; 80:1252-4.

35. Frazee RC, Roberts JW, Okeson GC, Symmonds RE, Snyder SK, Hendricks JC, Smith RW. Open versus laparoscopic cholecystectomy: a comparison of postoperative pulmonary function. Ann Surg 1991; 213:651-4.

36. Schwenk W, Bohm B, Muller JM. Postoperative pain and fatigue afterlaparoscopic or conventional colorectal resections: a prospective randomized trial.Surg Endosc 1998; 12: 1131–1136.